



## [0,1] Truncated Fréchet-Weibull and Fréchet Distributions

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### ABSTRACT

In this paper, we introduce a new family of continuous distributions based on [0, 1] Truncated Fréchet distribution. [0, 1] Truncated Fréchet Weibull ([0, 1]TFW) and [0, 1] Truncated Fréchet ([0, 1] TFF) distributions are discussed as special cases. The cumulative distribution function, the  $r$ th moment, the mean, the variance, the skewness, the kurtosis, the mode, the median, the characteristic function, the reliability function and the hazard rate function are obtained for the distributions under consideration. It is well known that an item fails when a stress to which it is subjected exceeds the corresponding strength. In this sense, strength can be viewed as “resistance to failure.” Good design practice is such that the strength is always greater than the expected stress. The safety factor can be defined in terms of strength and stress as strength/stress. So, the [0, 1] TFW strength-stress and the [0, 1] TFF strength-stress models with different parameters will be derived here. The Shannon entropy and Relative entropy will be derived also.

**Keywords:** [0, 1] TFW, [0, 1] TFF, Stress-strength model, Shannon entropy, Relative entropy.

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### 1. Introduction

Probability distribution analysis of lifetime data is an important topic in reliability, biomedical, finance, social sciences and others. Lately, efforts have been made to derive new classes of statistical distributions that expand famed families of distributions and then get major pliancy in modeling data in practical applications. Therefore, by appending new parameters to a deriving distribution we obtain classes of more flexible distributions. Here, we propose a distribution with the hope it will attract wider applicability in other fields. The generalization, which is motivated by the work of Eugene et al., will be our guide. Eugene et al. [2] defined the beta G distribution from a quite arbitrary Cumulative Distribution Function (CDF),  $G(x)$  by

$$F(x) = (1/\beta(a, b)) \int_0^{G(x)} w^{a-1} (1-w)^{b-1} dw, \quad (1)$$

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where  $a > 0$  and  $b > 0$  are two additional parameters whose role are to introduce skewness and to vary tail weight, and  $\beta(a, b) = \int_0^1 w^{a-1} (1-w)^{b-1} dw$  is the beta function. The class of distributions (1) has an increased attention after the works by Eugene et al. and Jones [2, 5]. Application of  $X = G^{-1}(V)$  to the random variable  $V$  following a beta distribution with parameters  $a$  and  $b$ ,  $V \sim B(a, b)$  say, and yields  $X$  with CDF (1). Eugene et al. [2] defined the Beta Normal (BN) distribution by taking  $G(x)$  to be the CDF of the normal distribution and derived some of its first moments. General expressions for the moments of the BN distribution were derived [4]. An extensive review of scientific literature on this subject is available in Abid and Hassan [1]. We can write (1) as

$$F(x) = I_{G(x)}(a, b), \quad (2)$$

where  $I_y(a, b) = (1/B(a, b)) \int_0^y w^{a-1} (1-w)^{b-1} dw$  denotes the incomplete beta function ratio, i.e. the CDF of the beta distribution with parameters  $a$  and  $b$ . For general  $a$  and  $b$ , we can express (2) in terms of the well-known hypergeometric function defined by,

$${}_2F_1(\alpha, \beta, \gamma; x) = \sum_{i=0}^{\infty} \frac{(\alpha)_i (\beta)_i}{(\gamma)_i i!} x^i,$$

where  $(\alpha)_i = \alpha(\alpha + 1) \dots (\alpha + i - 1)$  denotes the ascending factorial. We obtain

$$F(x) = \frac{G(x)^a}{a B(a, b)} {}_2F_1(a, 1-b, a+1; G(x)).$$

Based on the properties of the CDF,  $F(x)$  for any beta  $G$  distribution defined from a parent  $G(x)$  in (1), could, in principle, follow the properties of the hypergeometric function, which are well established in the literature (See Section 9.1 of Gradshteyn and Ryzhik [3]). The Probability Density Function (PDF) corresponding to (1) can be written in the form

$$f(x) = \frac{1}{B(a, b)} G(x)^{a-1} (1-G(x))^{b-1} g(x), \quad (3)$$

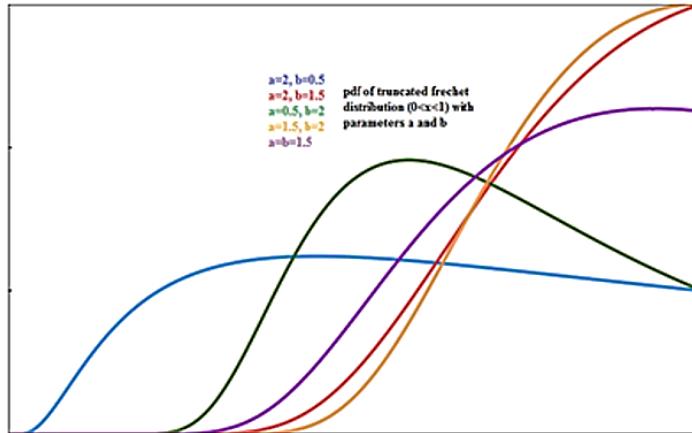
where  $g(x) = \partial G(x)/\partial x$  is the PDF of the parent distribution.

Now, the PDF and CDF of [0, 1] truncated Fréchet distribution are respectively

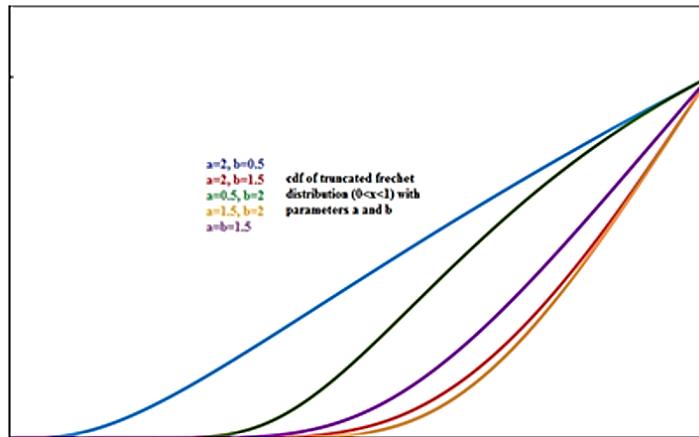
$$h(x) = \frac{ab}{e^{-a}} x^{-(b+1)} e^{-ax^{-b}} \quad 0 < x < 1, \quad (4)$$

$$H(x) = \frac{1}{e^{-a}} e^{-ax^{-b}}. \quad (5)$$

Graphs for some arbitrary parameters values of PDF and CDF are shown in Figure (1) and Figure (2), respectively



**Figure 1.** PDF of [0, 1] truncated Fréchet distribution.



**Figure 2.** CDF of [0, 1] truncated Fréchet distribution.

Now, Given two absolutely continuous CDFs , H and G, so that H and G are their corresponding PDFs. We suggest a new distribution F by composing H with G, so that  $F(x) = H(G(x))$  is a CDF:

$$\begin{aligned} F(x) &= \int_0^{G(x)} \frac{ab}{e^{-a}} t^{-(b+1)} e^{-at^{-b}} dt \\ &= \frac{1}{e^{-a}} e^{-at^{-b}} \Big|_0^{G(x)} = \frac{1}{e^{-a}} e^{-aG(x)^{-b}}, \end{aligned} \tag{6}$$

with PDF:

$$\begin{aligned} f(x) &= \frac{\partial}{\partial x} F(x) = \frac{\partial}{\partial x} \frac{e^{-aG(x)^{-b}}}{e^{-a}} \\ &= \frac{ab}{e^{-a}} e^{-aG(x)^{-b}} (G(x))^{-(b+1)} g(x). \end{aligned} \tag{7}$$

By considering  $G(x)$  as a baseline distribution, we define a generalized class of distributions in (6) and (7). We will name it the [0, 1] truncated Fréchet -G distribution. In the following two sections, we will assume that G are Weibull and Fréchet distributions, respectively.

## 2. [0, 1] Truncated Fréchet-Weibull Distribution

Assume that  $g(x) = k/\lambda (x/\lambda)^{k-1} \text{Exp}\{-(x/\lambda)^k\}$  and  $G(x) = 1 - \text{Exp}\{-(x/\lambda)^k\}$  ( $0 \leq x$ ) are PDF and CDF of Weibull random variable, respectively. Then, by applying (6) and (7), we get the CDF and PDF [0, 1] TFW random variable as follows:

$$F(x) = \frac{1}{e^{-a}} e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} \quad (8)$$

$$f(x) = \frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} \quad x \geq 0. \quad (9)$$

So, the reliability  $R(x)$  and hazard rate  $\lambda(x)$  functions are respectively

$$R(x) = 1 - \frac{e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}}}{e^{-a}} = 1 - e^{-a \left[\left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b} - 1\right]},$$

$$\lambda(x) = \frac{\frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}}}{1 - e^{-a \left[\left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b} - 1\right]}}.$$

The rth raw moment can be derived as follow:

$$E(x^r) = \int_0^\infty x^r \frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx$$

$$= \frac{abk}{\lambda e^{-a}} \int_0^\infty \lambda^{-(k-1)} x^r x^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx.$$

By using poison series,  $e^{-a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} = \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^m \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-bm}$ , we get

$$E(x^r) = \frac{abk}{\lambda e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^m \int_0^\infty \lambda^{-k+1} x^{r+k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-bm} dx$$

$$= \frac{bk}{\lambda^k e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \int_0^\infty x^{r+k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bm+b+1)} dx.$$

By using the series expansion,  $(1-z)^{-k} = \sum_{j=0}^{\infty} \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j$   $|z| < 1, k > 0$ , we get

$$\left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bm+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} e^{-j \left(\frac{x}{\lambda}\right)^k}, \text{ and then}$$

$$\begin{aligned}
E(x^r) &= \frac{bk}{\lambda^k e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \int_0^{\infty} x^{r+k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} e^{-j\left(\frac{x}{\lambda}\right)^k} dx \\
&= \frac{bk}{\lambda^k e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \int_0^{\infty} x^{r+k-1} e^{-\left(\frac{x}{\lambda}\right)^k} e^{-j\left(\frac{x}{\lambda}\right)^k} dx \\
E(x^r) &= \frac{bk}{\lambda^k e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \int_0^{\infty} x^{r+k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx.
\end{aligned}$$

Let  $y = \left(\frac{x}{\lambda}\right)^k \Rightarrow x = \lambda y^{\frac{1}{k}} \rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy$ , then

$$\begin{aligned}
E(x^r) &= \frac{bk}{\lambda^k e^{-a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \int_0^{\infty} \lambda^{r+k-1} y^{\frac{r}{k}+1-\frac{1}{k}} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\
&= b e^a \lambda^r \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \int_0^{\infty} y^{\frac{r}{k}} e^{-(j+1)y} dy \\
E(x^r) &= b e^a \lambda^r \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{r}{k}+1\right)}{(j+1)^{\frac{r}{k}+1}}. \tag{10}
\end{aligned}$$

Then, the characteristic function is

$$\begin{aligned}
Q_x(t) &= E(e^{ixt}) \\
&= \sum_{r=0}^{\infty} \frac{(it)^r}{r!} E(x^r), \text{ since } e^{ixt} = \sum_{r=0}^{\infty} \frac{(it)^r}{r!} x^r \\
&= b e^a \lambda^r \sum_{r=0}^{\infty} \frac{(it)^r}{r!} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{r}{k}+1\right)}{(j+1)^{\frac{r}{k}+1}} \\
&= b e^a \sum_{r=0}^{\infty} \frac{(it\lambda)^r}{r!} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{r}{k}+1\right)}{(j+1)^{\frac{r}{k}+1}}.
\end{aligned}$$

So, the mean  $\mu$  and variance  $\sigma^2$  of the of [0, 1] TFW random variable are

$$\begin{aligned}
\mu &= E(x) = b e^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{1}{k}+1\right)}{(j+1)^{\frac{1}{k}+1}}. \tag{11} \\
\sigma^2 &= E(x^2) - (Ex)^2 \\
&= b e^a \lambda^2 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{2}{k}+1\right)}{(j+1)^{\frac{2}{k}+1}} - b^2 e^{2a} \lambda^2 \\
&\quad \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{1}{k}+1\right)}{(j+1)^{\frac{1}{k}+1}} \right\}^2.
\end{aligned}$$

$$\sigma^2 = be^a \lambda^2 \left[ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} - be^a \right] \left[ \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \right]. \quad (12)$$

Since  $F(x) = \frac{e^{-a(1-e^{-\frac{x}{\lambda}})^k}}{e^{-a}} = \frac{1}{2}$ , the median  $M_e$  can be obtained as

$$x = M_e = \lambda \left\{ -\ln \left[ 1 - \left( 1 + \frac{\ln(2)}{a} \right)^{\frac{-1}{b}} \right] \right\}^{\frac{1}{k}}. \quad (13)$$

The skewness of [0, 1] TFW random variable will be  $sk = \frac{\mu_3}{(\mu_2)^{3/2}} = \frac{Ex^3 - 3\mu Ex^2 + 2\mu^3}{(\sigma^2)^{3/2}}$

$$= \frac{\left\{ \begin{array}{l} be^a \lambda^3 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{3}{k}+1)}{(j+1)^{\frac{3}{k}+1}} - 3 \\ \left\{ be^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\} \\ \left\{ be^a \lambda^2 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \right\} + \\ 2 \left\{ be^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^3 \end{array} \right\}}{\left\{ \begin{array}{l} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} - \\ be^a \left( \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right)^2 \end{array} \right\}^{3/2}} \quad (14)$$

In addition, the kurtosis is  $kr = \frac{\mu_3}{\mu_2^2} - 3 = \frac{Ex^4 - 4\mu Ex^3 + 6\mu^2 Ex^2 - 3\mu^4}{(\sigma^2)^2} - 3$

$$\begin{aligned}
 & \left\{ \begin{array}{l} be^a \lambda^4 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{4}{k}+1)}{(j+1)^{\frac{4}{k}+1}} \\ -4 \left\{ be^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\} \\ \left\{ be^a \lambda^3 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{3}{k}+1)}{(j+1)^{\frac{3}{k}+1}} \right\} \\ +6 \left\{ be^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \\ \left\{ be^a \lambda^2 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \right\} \\ -3 \left\{ be^a \lambda \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^4 \end{array} \right\} \\ = & \frac{\left\{ \begin{array}{l} be^a \lambda^2 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \\ -b^2 e^{2a} \lambda^2 \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \end{array} \right\}^2 - 3}{\left\{ \begin{array}{l} be^a \lambda^2 \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \\ -b^2 e^{2a} \lambda^2 \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \end{array} \right\}^2} - 3
 \end{aligned}$$

$$\begin{aligned}
& \left( \frac{1}{b^3 e^{3a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{4}{k}+1)}{(j+1)^{\frac{4}{k}+1}} \right. \\
& - \frac{4}{b^2 e^{2a}} \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\} \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{3}{k}+1)}{(j+1)^{\frac{3}{k}+1}} \right\} + \frac{6}{b e^a} \right\} \\
& b^4 e^{4a} \lambda^4 \left\{ \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \right\} - 3 \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^4 \right\} \\
& = \frac{b^4 e^{4a} \lambda^4}{\left( \frac{1}{b e^a} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \right)^2 - 3}
\end{aligned}$$

$$\begin{aligned}
& \left\{ \frac{1}{b^3 e^{3a}} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{4}{k}+1)}{(j+1)^{\frac{4}{k}+1}} - \right. \\
& \left. \frac{4}{b^2 e^{2a}} \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\} \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{3}{k}+1)}{(j+1)^{\frac{3}{k}+1}} \right\} + \frac{6}{b e^a} \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2 \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} \right\} - 3 \right. \\
& \left. \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^4 \right\} \\
= & \frac{\left( \frac{1}{b e^a} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{2}{k}+1)}{(j+1)^{\frac{2}{k}+1}} - \right)^2 - 3}{\left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma(\frac{1}{k}+1)}{(j+1)^{\frac{1}{k}+1}} \right\}^2} \quad (15)
\end{aligned}$$

The quintile function  $x_q$  of [0, 1] TFW random variable can be derived as

$$\begin{aligned}
q = P(x \leq x_q) = F_x(x_q) &= \frac{e^{-a \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b}}}{e^{-a}} \quad 0 < q < 1 \quad x_q > 0 \\
x_q = F^{-1}(q) &= \lambda \left\{ -\ln \left[ 1 - \left( 1 - \frac{\ln(q)}{a} \right)^{\frac{-1}{b}} \right] \right\}^{\frac{1}{k}}. \quad (16)
\end{aligned}$$

So, by using the inverse transform method, we can generate [0, 1] TFW random variable as follows:  $x = \lambda \left\{ -\ln \left[ 1 - \left( 1 - \frac{\ln(u)}{a} \right)^{\frac{-1}{b}} \right] \right\}^{\frac{1}{k}}$ , where u is a random number distribution uniformly in the unit interval [0, 1].

## 2.1. Shannon and Relative Entropies

An entropy of a random variable  $X$  is a measure of variation of the uncertainty. The Shannon entropy of  $[0, 1]$  TFW( $a, b, \lambda, k$ ) random variable  $X$  can be found as follows:

$$\begin{aligned} H &= - \int_0^\infty f(x) \ln(f(x)) dx \\ &= - \int_0^\infty f(x) \ln \left( \frac{abk}{\lambda e^{-a}} \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} \right) dx \\ &= \ln \left( \frac{\lambda e^{-a}}{abk} \right) - (k-1)E \left( \ln \left( \frac{x}{\lambda} \right) \right) + E \left( \left( \frac{x}{\lambda} \right)^k \right) + (b+1)E \left( \ln \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right) \right) \\ &\quad + aE \left( \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b} \right). \end{aligned}$$

$$\text{Let } I_1 = -(k-1)E \left( \ln \left( \frac{x}{\lambda} \right) \right),$$

$$H = -(k-1) \frac{abk}{\lambda e^{-a}} \int_0^\infty \ln \left( \frac{x}{\lambda} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} dx.$$

$$\text{Since } e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} = \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-bi}, \text{ then}$$

$$I_1 = -(k-1) \frac{abk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \int_0^\infty \ln \left( \frac{x}{\lambda} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-bi} dx$$

$$I_1 = -(k-1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^\infty \ln \left( \frac{x}{\lambda} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+b+1)} dx.$$

By using expansion series  $(1-z)^{-k} = \sum_{j=0}^{\infty} \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j$   $|z| < 1, k > 0$ , we get

$$\left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-[b(i+1)+1]} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j \left( \frac{x}{\lambda} \right)^k}, \text{ and then}$$

$$\begin{aligned} I_1 &= -(k-1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \ln \left( \frac{x}{\lambda} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} e^{-j \left( \frac{x}{\lambda} \right)^k} dx \\ &= -(k-1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \ln \left( \frac{x}{\lambda} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-(j+1) \left( \frac{x}{\lambda} \right)^k} dx. \end{aligned}$$

$$\text{Let } y = \left( \frac{x}{\lambda} \right)^k \Rightarrow x = \lambda y^{\frac{1}{k}} \Rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy, \text{ then}$$

$$\begin{aligned} I_1 &= -(k-1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \ln \left( y^{\frac{1}{k}} \right) \left( \frac{\lambda y^{\frac{1}{k}}}{\lambda} \right)^{k-1} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\ &= \frac{-(k-1)be^a}{k} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \ln(y) e^{-(j+1)y} dy. \end{aligned}$$

Since  $\int_0^\infty x^{s-1} e^{-mx} \ln(x) dx = m^{-s} \Gamma(s) \{\Psi(s) - \ln(m)\}$ , where

$\Psi(1) = -\ln(\gamma) \sim 0.5772$  and  $\gamma = 0.5772$ , then

$$\begin{aligned} I_1 &= \frac{(k-1)b e^a}{k} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)} \{\gamma + \ln(j+1)\}, \text{ and } I_2 = E \left( \left( \frac{x}{\lambda} \right)^k \right)^{-b} \\ &= \frac{abk}{\lambda e^{-a}} \int_0^\infty \left( \frac{x}{\lambda} \right)^k \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left( \frac{x}{\lambda} \right)^k} \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b}} dx \\ &= \frac{abk}{\lambda e^{-a}} \int_0^\infty \left( \frac{x}{\lambda} \right)^{2k-1} e^{-\left( \frac{x}{\lambda} \right)^k} \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b}} dx. \end{aligned}$$

Since  $e^{-a \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b}} = \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-bi}$ , then

$$I_2 = \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^\infty \left( \frac{x}{\lambda} \right)^{2k-1} e^{-\left( \frac{x}{\lambda} \right)^k} \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-(bi+b+1)} dx.$$

By using expansion series  $(1-z)^{-k} = \sum_{j=0}^{\infty} \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j$   $|z| < 1$ ,  $k > 0$ , we get

$$\left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-(bi+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j \left( \frac{x}{\lambda} \right)^k}, \text{ and then}$$

$$I_2 = \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \left( \frac{x}{\lambda} \right)^{2k-1} e^{-(j+1) \left( \frac{x}{\lambda} \right)^k} dx.$$

Let  $y = \left( \frac{x}{\lambda} \right)^k \rightarrow x = \lambda y^{\frac{1}{k}} \rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy$

$$\begin{aligned} I_2 &= \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \left( \frac{\lambda y^{\frac{1}{k}}}{\lambda} \right)^{2k-1} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\ &= b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty y e^{-(j+1)y} dy \\ &= b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)^2}, \end{aligned}$$

and  $I_3 = (b+1)E \left( \ln \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right) \right)$

$$= \frac{(b+1)abk}{\lambda e^{-a}} \int_0^\infty \ln \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left( \frac{x}{\lambda} \right)^k} \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b}} dx.$$

Let  $y = \left( 1 - e^{-\left( \frac{x}{\lambda} \right)^k} \right)^{-b} \Rightarrow x = \lambda \left[ -\ln \left( 1 - y^{\frac{1}{b}} \right) \right]^{\frac{1}{k}}$

$$\begin{aligned}
&\Rightarrow dx = \frac{\lambda}{k} \left[ -\ln \left( 1 - y^{\frac{-1}{b}} \right) \right]^{\frac{1}{k}-1} \frac{-1}{\left( 1 - y^{\frac{-1}{b}} \right) b} y^{\frac{-1}{b}-1} dy \\
&\Rightarrow dx = \frac{\lambda}{k} \left[ -\ln \left( 1 - y^{\frac{-1}{b}} \right) \right]^{\frac{1}{k}-1} \frac{-1}{\left( 1 - y^{\frac{-1}{b}} \right) b} y^{\frac{-1}{b}-1} dy \\
I_3 &= \frac{(b+1)abk}{\lambda e^{-a}} \int_1^\infty \ln \left( y^{\frac{-1}{b}} \right) \left( \lambda \left[ -\ln \left( 1 - y^{\frac{-1}{b}} \right) \right]^{\frac{1}{k}} / \lambda \right)^{k-1} \left( 1 - y^{\frac{-1}{b}} \right) \left( y^{\frac{-1}{b}} \right)^{-(b+1)} e^{-ay} \\
&\quad \frac{\lambda}{k} \left[ -\ln \left( 1 - y^{\frac{-1}{b}} \right) \right]^{\frac{1}{k}-1} \frac{y^{\frac{-1}{b}-1}}{b \left( 1 - y^{\frac{-1}{b}} \right)} dy \\
&= \frac{-(b+1)ae^a}{b} \int_1^\infty \ln(y) e^{-ay} dy \\
&= \frac{-(b+1)ae^a}{b} \left[ \int_0^\infty \ln(y) e^{-ay} dy - \int_0^1 \ln(y) e^{-ay} dy \right] \\
I_{31} &= \int_0^\infty \ln(y) e^{-ay} dy = \frac{-1}{a} \{ \gamma + \ln(a) \}.
\end{aligned}$$

Since  $\int_0^\infty x^{s-1} e^{-mx} \ln(x) dx = m^{-s} \Gamma(s) \{ \Psi(s) - \ln(m) \}$ , where  $\Psi(1) = -\gamma$ ,  $\gamma = 0.57721$  is an Eualer constant.

$$\begin{aligned}
I_{32} &= \int_0^1 \ln(y) e^{-ay} dy = \sum_{m=0}^\infty \frac{(-a)^m}{m!} \int_0^1 y^m \ln(y) dy = - \sum_{m=0}^\infty \frac{(-a)^m}{m!(m+1)^2}, \text{ since} \\
e^{-ay} &= \sum_{m=0}^\infty \frac{(-a)^m}{m!} y^m \text{ and } \int x^m \ln(x) dx = x^{m+1} \left\{ \frac{\ln(x)}{m+1} - \frac{1}{(m+1)^2} \right\}, \\
I_3 &= \frac{(b+1)e^a}{b} \left\{ \gamma + \ln(a) - \sum_{m=0}^\infty \frac{(-1)^m a^{m+1}}{m!(m+1)^2} \right\}, \text{ and } I_4 = a E \left( \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b} \right) \\
&= \frac{a^2 bk}{\lambda e^{-a}} \int_0^\infty \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} dx.
\end{aligned}$$

Since  $e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} = \sum_{i=0}^\infty \frac{(-1)^i}{i!} a^i \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-bi}$ , then

$$I_4 = \frac{bk}{\lambda e^{-a}} \sum_{i=0}^\infty \frac{(-1)^i}{i!} a^{i+2} \int_0^\infty \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+2b+1)} dx.$$

By using expansion series  $(1-z)^{-k} = \sum_{j=0}^\infty \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j$   $|z| < 1$ ,  $k > 0$ , we get

$$\begin{aligned}
&\left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+2b+1)} = \sum_{j=0}^\infty \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} e^{-j \left(\frac{x}{\lambda}\right)^k}, \text{ then} \\
I_4 &= \frac{bk}{\lambda e^{-a}} \sum_{i=0}^\infty \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^\infty \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \int_0^\infty \left(\frac{x}{\lambda}\right)^{k-1} e^{-(j+1) \left(\frac{x}{\lambda}\right)^k} dx.
\end{aligned}$$

Let  $y = \left(\frac{x}{\lambda}\right)^k \rightarrow x = \lambda y^{\frac{1}{k}} \rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy$

$$\begin{aligned}
 I_4 &= \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \int_0^{\infty} \left(y^{\frac{1}{k}}\right)^{k-1} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\
 &= be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \int_0^{\infty} e^{-(j+1)y} dy \\
 &= be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \frac{1}{(j+1)} \\
 H &= \ln\left(\frac{\lambda e^{-a}}{abk}\right) + \frac{(k-1)}{k} be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)} \{y \\
 &\quad + \ln(j+1)\} \\
 &\quad + be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)^2} + \frac{(b+1)}{b} e^a \\
 &\left\{y + \ln(a) - \sum_{m=0}^{\infty} \frac{(-1)^m a^{m+1}}{m! (m+1)^2}\right\} + b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \frac{1}{(j+1)}. \tag{17}
 \end{aligned}$$

The relative entropy or the Kullback–Leibler divergence is a measure of the difference between two probability distributions  $F_1$  and  $F_2$ . It is not symmetric in  $F_1$  and  $F_2$ . In applications,  $F_1$  typically represents the “true” distribution of data, observations, or a precisely calculated theoretical distribution, while  $F_2$  typically represents a theory, model, description, or approximation of  $F_1$ . Specifically, the Kullback–Leibler divergence of  $F_2$  from  $F_1$ , denoted  $D_{KL}(F_1||F_2)$ , is a measure of the information gained when one revises beliefs from the prior probability distribution  $F_2$  to the posterior probability distribution  $F_1$ . More exactly, it is the amount of information that is *lost* when  $F_2$  is used to approximate  $F_1$ , defined operationally as the expected extra number of bits required to code samples from  $F_1$  using a code optimized for  $F_2$  rather than the code optimized for  $F_1$ . So, the relative entropy  $D_{KL}(F||F^*)$  for a random variable  $[0, 1]$  TFW( $a, b, \lambda, k$ ) can be found as follows:

$$\frac{f(x)}{f^*(x)} = \frac{abk\lambda_1 e^{-a} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}}}{\alpha\beta k_1 \lambda e^{-a} \left(\frac{x}{\lambda_1}\right)^{k_1-1} e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-(\beta+1)} e^{-\alpha\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-\beta}}}$$

$$\begin{aligned}
& Dkl(F \| F^*) \\
&= \int_0^\infty f(x) \ln \left( \frac{abk\lambda_1 e^{-\alpha} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}}}{\alpha\beta k_1 \lambda e^{-a} \left(\frac{x}{\lambda_1}\right)^{k_1-1} e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-(\beta+1)} e^{-\alpha\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-\beta}}} \right) dx \\
&= \int_0^\infty f_1(x) \left[ a \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b} - (k_1 - 1) \ln \left(\frac{x}{\lambda_1}\right) + \left(\frac{x}{\lambda_1}\right)^{k_1} + (\beta + 1) \ln \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right) \right. \\
&\quad \left. + \alpha \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-\beta} \right] dx \\
&= \ln \left( \frac{abk\lambda_1 e^{-\alpha}}{\alpha\beta k_1 \lambda e^{-a}} \right) + (k - 1)E \left( \ln \left(\frac{x}{\lambda}\right) \right) - E \left( \left(\frac{x}{\lambda}\right)^k \right) - (b + 1)E \left( \ln \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right) \right) - a \\
&E \left( \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b} \right) - (k_1 - 1)E \left( \ln \left(\frac{x}{\lambda_1}\right) \right) + E \left( \left(\frac{x}{\lambda_1}\right)^{k_1} \right) + (\beta + 1)E \left( \ln \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right) \right) + \\
&\quad \alpha E \left( \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-\beta} \right).
\end{aligned}$$

Let,  $I_1 = (k - 1)E \left( \ln \left(\frac{x}{\lambda}\right) \right)$

$$\begin{aligned}
&= (k - 1) \frac{abk}{\lambda e^{-a}} \int_0^\infty \ln \left(\frac{x}{\lambda}\right) \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx \\
&= -\frac{(k - 1)}{k} b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)} \{Y + \ln(j+1)\},
\end{aligned}$$

and  $I_2 = -E \left( \left(\frac{x}{\lambda}\right)^k \right)$

$$\begin{aligned}
&= -\frac{abk}{\lambda e^{-a}} \int_0^\infty \left(\frac{x}{\lambda}\right)^k \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx \\
&= -b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)^2},
\end{aligned}$$

and  $I_3 = -(b + 1)E \left( \ln \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right) \right)$

$$= -(b + 1) \int_0^\infty \ln \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right) \frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx$$

$$= \frac{(b+1)}{b} e^a \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} \frac{a^{m+1}}{(m+1)^2} - (\mathbb{Y} + \ln(a)) \right\}$$

$$\begin{aligned} \text{and } I_4 &= -a E \left( \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b} \right) \\ &= -a \int_0^{\infty} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b} \frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} dx \\ &= -b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b(i+2)+1])} \frac{1}{(j+1)}, \end{aligned}$$

$$\begin{aligned} \text{and } I_5 &= -(k_1 - 1) E \left( \ln \left( \frac{x}{\lambda_1} \right) \right) \\ &= -(k_1 - 1) \frac{abk}{\lambda e^{-a}} \int_0^{\infty} \ln \left( \frac{x}{\lambda_1} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} dx. \end{aligned}$$

Since  $e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} = \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-bi}$ , then

$$I_5 = -(k_1 - 1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^{\infty} \ln \left( \frac{x}{\lambda_1} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+b+1)} dx.$$

By using expansion series  $(1-z)^{-k} = \sum_{j=0}^{\infty} \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j$   $|z| < 1, k > 0$ , we get

$$\left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j \left(\frac{x}{\lambda}\right)^k}, \text{ and then}$$

$$I_5 = -(k_1 - 1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^{\infty} \ln \left( \frac{x}{\lambda_1} \right) \left( \frac{x}{\lambda} \right)^{k-1} e^{-(j+1) \left(\frac{x}{\lambda}\right)^k} dx.$$

$$\text{Let } y = \left(\frac{x}{\lambda}\right)^k \rightarrow x = \lambda y^{\frac{1}{k}} \rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy$$

$$\begin{aligned} I_5 &= -(k_1 - 1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^{\infty} \ln \left( \frac{\lambda y^{\frac{1}{k}}}{\lambda_1} \right) \left( y^{\frac{1}{k}} \right)^{k-1} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\ &= -(k_1 - 1) b e^a \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^{\infty} \left[ \ln \left( \frac{\lambda}{\lambda_1} \right) + \frac{1}{k} \ln y \right] e^{-(j+1)y} dy \\ &= (k_1 - 1) b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left\{ \frac{1}{k(j+1)} (\mathbb{Y} + \ln(j+1)) \right. \\ &\quad \left. - \ln \left( \frac{\lambda}{\lambda_1} \right) \frac{1}{(j+1)} \right\}, \end{aligned}$$

$$\begin{aligned} \text{and } I_6 &= E\left(\left(\frac{x}{\lambda_1}\right)^{k_1}\right) = \left(\frac{1}{\lambda_1}\right)^{k_1} E(x^{k_1}) \\ &= \left(\frac{1}{\lambda_1}\right)^{k_1} b e^a \lambda^{k_1} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{k_1}{k}+1\right)}{(j+1)^{\frac{k_1}{k}+1}} \\ &= b e^a \left(\frac{\lambda}{\lambda_1}\right)^{k_1} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{k_1}{k}+1\right)}{(j+1)^{\frac{k_1}{k}+1}}, \end{aligned}$$

$$\begin{aligned} \text{and } I_7 &= (\beta+1)E\left(\ln\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)\right) \\ &= (\beta+1) \frac{abk}{\lambda e^{-a}} \int_0^\infty \ln\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right) \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} dx. \end{aligned}$$

$$\text{Since } e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} = \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-bi}, \text{ then}$$

$$\begin{aligned} I_7 &= (\beta+1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^\infty \ln\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right) \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bi+b+1)} dx \\ &\quad \left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bi+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j\left(\frac{x}{\lambda}\right)^k}, \text{ and then} \end{aligned}$$

$$\begin{aligned} I_7 &= (\beta+1) \frac{bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \left(\frac{x}{\lambda}\right)^{k-1} \ln\left(1-e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right) e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx \\ &\Rightarrow \text{let } y = 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \Rightarrow x = \lambda_1 [-\ln(1-y)]^{\frac{1}{k_1}} \Rightarrow dx = \frac{\lambda_1}{k_1} [-\ln(1-y)]^{\frac{1}{k_1}-1} \frac{1}{(1-y)} dy, \end{aligned}$$

then

$$\begin{aligned} I_7 &= \frac{(\beta+1)bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^1 \ln(y) \left(\frac{\lambda_1 [-\ln(1-y)]^{\frac{1}{k_1}}}{\lambda}\right)^{k-1} \\ &\quad e^{-(j+1)\left(\frac{\lambda_1 [-\ln(1-y)]^{\frac{1}{k_1}}}{\lambda}\right)^k} \frac{\lambda_1}{k_1} [-\ln(1-y)]^{\frac{1}{k_1}-1} \frac{1}{1-y} dy \\ &= \frac{(\beta+1)bk}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^{k-1} \left(\frac{\lambda_1}{k_1}\right) \int_0^1 \ln(y) [-\ln(1-y)]^{\frac{k}{k_1}-1} \\ &\quad e^{-(j+1)\left(\frac{\lambda_1}{\lambda}\right)^k [-\ln(1-y)]^{\frac{k}{k_1}}} \frac{1}{(1-y)} dy. \end{aligned}$$

Since  $e^{-(j+1)\left(\frac{\lambda_1}{\lambda}\right)^k[-\ln(1-y)]^{\frac{k}{k_1}}} = \sum_{v=0}^{\infty} \frac{(-1)^v}{v!} \left[(j+1)\left(\frac{\lambda_1}{\lambda}\right)^k\right]^v [-\ln(1-y)]^{\frac{vk}{k_1}}$ , then

$$\begin{aligned} I_7 &= (\beta + 1) \frac{bk}{\lambda e^{-a}} \sum_{i,v=0}^{\infty} \frac{(-1)^{i+v}}{i! v!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^k \left(\frac{\lambda}{\lambda_1}\right) \left(j+1\right) \left(\frac{\lambda_1}{\lambda}\right)^k \\ &\quad \int_0^1 \ln(y) [-\ln(1-y)]^{\frac{k}{k_1}-1} [-\ln(1-y)]^{\frac{vk}{k_1}} (1-y)^{-1} dy \\ &= (\beta + 1) \frac{bk}{k_1 e^{-a}} \sum_{i,v=0}^{\infty} \frac{(-1)^{i+v}}{i! v!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^k \left[j+1\right] \left(\frac{\lambda_1}{\lambda}\right)^k \int_0^1 \ln(y) \\ &\quad [-\ln(1-y)]^{\frac{vk}{k_1}+\frac{k}{k_1}-1} (1-y)^{-1} dy. \end{aligned}$$

Since the expansion of  $[-\ln(1-y)]^{\frac{vk}{k_1}+\frac{k}{k_1}-1} =$

$$\left(\frac{vk}{k_1} + \frac{k}{k_1} - 1\right) \sum_{m=0}^{\infty} \binom{m+1 - \frac{vk}{k_1} - \frac{k}{k_1}}{m} \sum_{u=0}^m \frac{(-1)^{m+u}}{\frac{vk}{k_1} + \frac{k}{k_1} - 1 - u} \binom{m}{u} p_{u,m} y^{\frac{vk}{k_1} + \frac{k}{k_1} + m - 1},$$

where  $p_{u,m}$  is constant can be calculated as

$$p_{u,m} = \frac{1}{m} \sum_{n=1}^m \frac{(-1)^n [n(u+1)-m]}{(n+1)} p_{u,m-n}, \text{ for } m = 1, 2, 3, \dots \text{ and } p_{u,0} = 1.$$

By using  $(a+b)^n = a^n + na^{n-1}b + \frac{n(n-1)}{2!}a^{n-2}b^2 + \frac{n(n-1)(n-2)}{3!}a^{n-3}b^3 + \dots$ , we get

$(1-y)^{-1} = 1 + y + y^2 + y^3 + \dots = \sum_{r=0}^{\infty} y^r$ , and then

$$\begin{aligned} I_7 &= \frac{(\beta + 1)bk}{k_1 e^{-a}} \sum_{i,v,r=0}^{\infty} \frac{(-1)^{i+v}}{i! v!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^k \left[j+1\right] \left(\frac{\lambda_1}{\lambda}\right)^k \\ &\quad \left(\frac{vk}{k_1} + \frac{k}{k_1}\right. \\ &\quad \left.- 1\right) \sum_{m=0}^{\infty} \binom{m+1 - \frac{vk}{k_1} - \frac{k}{k_1}}{m} \sum_{u=0}^m \frac{(-1)^{m+u}}{\frac{vk}{k_1} + \frac{k}{k_1} - 1 - u} \binom{m}{u} p_{u,m} \int_0^1 \ln(y) y^{\frac{vk}{k_1} + \frac{k}{k_1} + m + r - 1} dy. \end{aligned}$$

Since  $\int_0^1 \left(\ln\left(\frac{1}{y}\right)\right)^{\theta-1} y^{v-1} dy = \frac{1}{v^\theta} \Gamma(\theta)$ , then

$$\int_0^1 \ln(y) y^{\frac{vk}{k_1} + \frac{k}{k_1} + m + r - 1} dy = - \int_0^1 \ln\left(\frac{1}{y}\right) y^{\frac{vk}{k_1} + \frac{k}{k_1} + m + r - 1} dy = \frac{1}{\left(\frac{vk}{k_1} + \frac{k}{k_1} + m + r\right)^2} \Gamma(2), \text{ so}$$

$$\begin{aligned} I_7 &= \frac{(\beta + 1)bk}{k_1 e^{-a}} \sum_{i,v,r=0}^{\infty} \frac{(-1)^{i+v+1}}{i! v!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^k \left[j+1\right] \left(\frac{\lambda_1}{\lambda}\right)^k \\ &\quad \left(\frac{vk}{k_1} + \frac{k}{k_1} - 1\right) \sum_{m=0}^{\infty} \binom{m+1 - \frac{vk}{k_1} - \frac{k}{k_1}}{m} \sum_{u=0}^m \frac{(-1)^{m+u}}{\frac{vk}{k_1} + \frac{k}{k_1} - 1 - u} \binom{m}{u} p_{u,m} \frac{1}{\left(\frac{vk}{k_1} + \frac{k}{k_1} + m + r\right)^2}, \end{aligned}$$

$$\text{and } I_8 = \alpha E \left( \left( 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \right)^{-\beta} \right)$$

$$I_8 = \frac{\alpha abk}{\lambda e^{-a}} \int_0^\infty \left( 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \right)^{-\beta} \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(b+1)} e^{-a \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-b}} dx$$

$$= \frac{\alpha b k}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^\infty \left( 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \right)^{-\beta} \left( \frac{x}{\lambda} \right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+b+1)} dx.$$

Since  $\left( 1 - e^{-\left(\frac{x}{\lambda}\right)^k} \right)^{-(bi+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j\left(\frac{x}{\lambda}\right)^k}$ , then

$$I_8 = \frac{\alpha b k}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^\infty \left( 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \right)^{-\beta} \left( \frac{x}{\lambda} \right)^{k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx.$$

By using  $\left( 1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}} \right)^{-\beta} = \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} e^{-m\left(\frac{x}{\lambda_1}\right)^{k_1}}$ , we get

$$I_8 = \frac{\alpha b k}{\lambda e^{-a}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \int_0^\infty e^{-m\left(\frac{x}{\lambda_1}\right)^{k_1}} \left( \frac{x}{\lambda} \right)^{k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx.$$

So, by using  $e^{-m\left(\frac{x}{\lambda_1}\right)^{k_1}} = \sum_{l=0}^{\infty} \frac{(-1)^l}{l!} m^l \left( \frac{x}{\lambda_1} \right)^{lk_1}$ , we get

$$I_8 = \frac{\alpha b k}{\lambda e^{-a}} \sum_{i=0}^{\infty} \sum_{l=0}^{\infty} \frac{(-1)^{i+l}}{i! l!} a^{i+1} m^l \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \int_0^\infty \left( \frac{x}{\lambda_1} \right)^{lk_1} \left( \frac{x}{\lambda} \right)^{k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx$$

$$= \frac{\alpha b k e^a}{\lambda_1^{lk_1} \lambda^k} \sum_{i=0}^{\infty} \sum_{l=0}^{\infty} \frac{(-1)^{i+l}}{i! l!} a^{i+1} m^l \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \int_0^\infty x^{lk_1+k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx$$

$$\Rightarrow \text{let } y = \left( \frac{x}{\lambda} \right)^k \Rightarrow x = \lambda y^{\frac{1}{k}} \Rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy, \text{ then}$$

$$I_8 = \frac{\alpha b e^a \lambda^{lk_1}}{\lambda_1^{lk_1}} \sum_{i=0}^{\infty} \sum_{l=0}^{\infty} \frac{(-1)^{i+l}}{i! l!} a^{i+1} m^l \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \int_0^\infty y^{\frac{lk_1}{k}} e^{-(j+1)y} dy$$

$$= \alpha b e^a \left( \lambda / \lambda_1 \right)^{lk_1} \sum_{i=0}^{\infty} \sum_{l=0}^{\infty} \frac{(-1)^{i+l}}{i! l!} a^{i+1} m^l \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \frac{\Gamma\left(\frac{lk_1}{k} + 1\right)}{(j+1)^{\frac{lk_1}{k}+1}}.$$

Then, the relative entropy is

$$Dkl(F_1 \| F_2) = \ln \left( \frac{\alpha b k \lambda_1 e^{-\alpha}}{\alpha \beta k_1 \lambda e^{-a}} \right) - \frac{(k-1)}{k} b e^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])}$$

$$\begin{aligned}
& \frac{1}{(j+1)} \{Y + \ln(j+1)\} - be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \frac{1}{(j+1)^2} \\
& + \frac{(b+1)}{b} e^a \left\{ \sum_{m=0}^{\infty} \frac{(-1)^m a^{m+1}}{m! (m+1)^2} - (Y + \ln(a)) \right\} \\
& - be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+2} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+2)+1]+j)}{j! \Gamma([b[i+2]+1])} \\
& \frac{1}{(j+1)} + (k_1 - 1) be^a \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left\{ \frac{1}{k(j+1)} (Y + \ln(j+1)) \right. \\
& \left. - \ln\left(\frac{\lambda}{\lambda_1}\right) \frac{1}{(j+1)} \right\} \\
& + be^a \left(\frac{\lambda}{\lambda_1}\right)^{k_1} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(m+1)+1]+j)}{j! \Gamma([b(m+1)+1])} \frac{\Gamma\left(\frac{k_1}{k}+1\right)}{(j+1)^{\frac{k_1}{k}+1}} + (\beta+1) \frac{bk}{k_1 e^{-a}} \\
& \sum_{i,v,r=0}^{\infty} \frac{(-1)^{i+v+1}}{i! v!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \left(\frac{\lambda_1}{\lambda}\right)^k \left[ (j+1) \left(\frac{\lambda_1}{\lambda}\right)^k \right]^v \\
& \left(\frac{vk}{k_1} + \frac{k}{k_1} - 1\right) \sum_{m=0}^{\infty} \binom{m+1 - \frac{vk}{k_1} - \frac{k}{k_1}}{m} \sum_{u=0}^m \frac{(-1)^{m+u}}{\frac{vk}{k_1} + \frac{k}{k_1} - 1 - u} \binom{m}{u} p_{u,m} \frac{1}{\left(\frac{vk}{k_1} + \frac{k}{k_1} + m + r\right)^2} \\
& + \alpha b e^a \left(\frac{\lambda}{\lambda_1}\right)^{lk_1} \sum_{i=0}^{\infty} \sum_{l=0}^{\infty} \frac{(-1)^{i+l}}{i! l!} a^{i+1} m^l \sum_{j=0}^{\infty} \frac{\Gamma([b(j+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{m=0}^{\infty} \frac{\Gamma(\beta+m)}{m! \Gamma(\beta)} \frac{\Gamma\left(\frac{lk_1}{k}+1\right)}{(j+1)^{\frac{lk_1}{k}+1}} \quad (18)
\end{aligned}$$

## 2.2. Stress –Strength Reliability

Let Y and X be the stress and strength random variables, independent of each other, follow [0,1] TFW( $a, b, \lambda, k$ ) and [0,1]TFW( $\alpha, \beta, \lambda_1, k_1$ ), respectively, then

$$\begin{aligned}
R &= P(y < x) = \int_0^\infty f_X(x) F_Y(x) dx \\
&= \int_0^\infty \frac{abk}{\lambda e^{-a}} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} \frac{e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta}}}{e^{-\alpha}} dx. \\
\text{Since } e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-b}} &= \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-bi}, \text{ then}
\end{aligned}$$

$$\begin{aligned}
R &= \frac{abk}{\lambda e^{-a} e^{-\alpha}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^i \int_0^\infty \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(b+1)} \left(1 - \right. \\
&\quad \left. e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-bi} e^{-a\left(1-e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta}} dx
\end{aligned}$$

$$= \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \int_0^{\infty} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bi+b+1)} e^{-\alpha \left(1 - e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta}} dx.$$

By using  $\left(1 - e^{-\left(\frac{x}{\lambda}\right)^k}\right)^{-(bi+b+1)} = \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} e^{-j\left(\frac{x}{\lambda}\right)^k}$ , we get

$$R = \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \int_0^{\infty} \left(\frac{x}{\lambda}\right)^{k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} e^{-\alpha \left(1 - e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta}} dx.$$

So, by using  $e^{-\alpha \left(1 - e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta}} = \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} \alpha^m \left(1 - e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta m}$ , we get

$$\begin{aligned} R &= \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \frac{(-1)^{i+m}}{i! m!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)]+1)} \sum_{m=0}^{\infty} \alpha^m \int_0^{\infty} \left(\frac{x}{\lambda}\right)^{k-1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} \\ &\quad \left(1 - e^{-\left(\frac{x}{\lambda}\right)^{k_1}}\right)^{-\beta m} dx \\ &= \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \sum_{m=0}^{\infty} \frac{(-1)^{i+m}}{i! m!} a^{i+1} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \alpha^m \sum_{u=0}^{\infty} \frac{\Gamma(\beta m + u)}{u! \Gamma(\beta m)} \int_0^{\infty} \left(\frac{x}{\lambda}\right)^{k-1} \\ &\quad e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} e^{-u\left(\frac{x}{\lambda}\right)^{k_1}} dx. \end{aligned}$$

Since  $e^{-u\left(\frac{x}{\lambda}\right)^{k_1}} = \sum_{v=0}^{\infty} \frac{(-1)^v}{v!} u^v \left(\frac{x}{\lambda}\right)^{vk_1}$ , then

$$\begin{aligned} R &= \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \sum_{m=0}^{\infty} \sum_{v=0}^{\infty} \frac{(-1)^{i+m+v}}{i! m! v!} a^{i+1} \alpha^m u^v \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{u=0}^{\infty} \frac{\Gamma(\beta m + u)}{u! \Gamma(\beta m)} \\ &\quad \int_0^{\infty} \lambda_1^{-vk_1} \left(\frac{x}{\lambda}\right)^{k-1} x^{vk_1} e^{-(j+1)\left(\frac{x}{\lambda}\right)^k} dx. \end{aligned}$$

Let  $y = \left(\frac{x}{\lambda}\right)^k \rightarrow x = \lambda y^{\frac{1}{k}} \rightarrow dx = \frac{\lambda}{k} y^{\frac{1}{k}-1} dy$ , then

$$\begin{aligned} R &= \frac{bk}{\lambda e^{-\alpha} e^{-\alpha}} \sum_{i=0}^{\infty} \sum_{m=0}^{\infty} \sum_{v=0}^{\infty} \frac{(-1)^{i+m+v}}{i! m! v!} a^{i+1} \alpha^m u^v \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \sum_{u=0}^{\infty} \frac{\Gamma(\beta m + u)}{u! \Gamma(\beta m)} \\ &\quad \int_0^{\infty} \lambda_1^{-vk_1} \left(y^{\frac{1}{k}}\right)^{k-1} \lambda^{vk_1} y^{\frac{vk_1}{k}} e^{-(j+1)y} \frac{\lambda}{k} y^{\frac{1}{k}-1} dy \\ &= b e^{\alpha+\alpha} \left(\frac{\lambda}{\lambda_1}\right)^{vk_1} \sum_{i=0}^{\infty} \sum_{m=0}^{\infty} \sum_{v=0}^{\infty} \frac{(-1)^{i+m+v}}{i! m! v!} a^{i+1} \alpha^m u^v \sum_{u=0}^{\infty} \frac{\Gamma(\beta m + u)}{u! \Gamma(\beta u)} \sum_{j=0}^{\infty} \frac{\Gamma([b(i+1)+1]+j)}{j! \Gamma([b(i+1)+1])} \\ &\quad \frac{\Gamma\left(\frac{vk_1}{k} + 1\right)}{(j+1)^{\frac{vk_1}{k}+1}}, \end{aligned} \tag{19}$$

$$\text{where } \left(1 - e^{-\left(\frac{x}{\lambda_1}\right)^{k_1}}\right)^{-\beta m} = \sum_{u=0}^{\infty} \frac{\Gamma(\beta m+u)}{u! \Gamma(\beta m)} e^{-u\left(\frac{x}{\lambda_1}\right)^{k_1}}.$$

### 3. [0, 1] Truncated Fréchet- Fréchet Distribution

We assume that  $g(x) = \alpha\beta x^{-(\beta+1)} \text{Exp}\{-\alpha x^{-\beta}\}$  and  $G(x) = \text{Exp}\{-\alpha x^{-\beta}\}$  ( $0 < x$ ) are PDF and CDF of Fréchet random variable, respectively. Then, by applying the equations (6) and (7), we get the PDF and CDF of  $[0, 1]$  TFF random variable as follows:

$$F(x) = \frac{e^{-ae^{\alpha bx^{-\beta}}}}{e^{-a}} \quad (20)$$

$$\begin{aligned} f(x) &= \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} \left(e^{-\alpha x^{-\beta}}\right)^{-b} e^{-a(e^{-\alpha x^{-\beta}})^{-b}} & x \geq 0 \\ f(x) &= \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} & x \geq 0. \end{aligned} \quad (21)$$

So, the reliability  $R(x)$  and hazard rate  $\lambda(x)$  functions are respectively

$$\begin{aligned} R(x) &= 1 - \frac{e^{-ae^{\alpha bx^{-\beta}}}}{e^{-a}} \\ &= 1 - e^{-a(e^{\alpha bx^{-\beta}} - 1)} \\ \lambda(x) &= \frac{f_1(x)}{R(x)} = \frac{ab\alpha\beta e^a x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}}}{1 - e^{-a(e^{\alpha bx^{-\beta}} - 1)}}. \end{aligned}$$

The rth raw moment can be derived as follows:

$$E(x^r) = \int_0^\infty x^r \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx.$$

Let  $y = e^{\alpha bx^{-\beta}} \Rightarrow x = \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} \frac{1}{ab} dy$ , then

$$\begin{aligned} E(x^r) &= \frac{ab\alpha\beta}{e^{-a}} \int_1^\infty \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} y e^{-ay} \frac{1}{\beta} \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} \frac{1}{aby} dy \\ &= ae^a \int_1^\infty \left(\frac{\ln y}{ab}\right)^{\frac{-r}{\beta}+1+\frac{1}{\beta}} e^{-ay} \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} dy \\ &= ae^a (\alpha b)^{\frac{r}{\beta}} \int_1^\infty (\ln y)^{\frac{-r}{\beta}} e^{-ay} dy \\ E(x^r) &= ae^a (\alpha b)^{\frac{r}{\beta}} \tau_u\left(a, \frac{r}{\beta}, 1\right), \end{aligned} \quad (22)$$

where  $\int_0^\infty (\ln y)^{\frac{-r}{\beta}} e^{-ay} dy = \tau\left(a, \frac{r}{\beta}\right)$  and  $\int_0^1 (\ln y)^{\frac{-r}{\beta}} e^{-ay} dy = \tau_l\left(a, \frac{r}{\beta}, 1\right)$

$$E(x^r) = ae^a(\alpha b)^{\frac{1}{\beta}} \left[ \int_0^\infty (\ln y)^{\frac{-r}{\beta}} e^{-ay} dy - \int_0^1 (\ln y)^{\frac{-r}{\beta}} e^{-ay} dy \right]$$

$$E(x^r) = ae^a(\alpha b)^{\frac{1}{\beta}} \left\{ \tau \left( a, \frac{r}{\beta} \right) - \tau_l \left( a, \frac{r}{\beta}, 1 \right) \right\},$$

and the characteristic function is

$$Q_x(t) = E(e^{itx})$$

$$= \sum_{r=0}^{\infty} \frac{(it)^r}{r!} E(x^r), \text{ since } e^{ixt} = \sum_{r=0}^{\infty} \frac{(it)^r}{r!} x^r$$

$$Q_x(t) = ae^a \sum_{r=0}^{\infty} \frac{(it(\alpha b)^{1/\beta})^r}{r!} \tau_u \left( a, \frac{r}{\beta}, 1 \right).$$

So, the mean  $\mu$  and variance  $\sigma^2$  of the [0, 1] TFF random variable are

$$\mu = E(x) = ae^a(\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \quad (23)$$

$$\sigma^2 = E(x^2) - (Ex)^2$$

$$\sigma^2 = a^2 e^{2a} (\alpha b)^{\frac{2}{\beta}} \left\{ \frac{1}{ae^a} \tau_u \left( a, \frac{2}{\beta}, 1 \right) - \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \right\}. \quad (24)$$

Since  $F(x) = \frac{e^{-ae^a\alpha bx^{-\beta}}}{e^{-a}} = \frac{1}{2}$ , then the median  $M_e$  can be calculated as

$$x = M_e = \left( \frac{\ln(1 + \ln(2)/a)}{\alpha b} \right)^{\frac{-1}{\beta}} \quad (25)$$

$$sk = \frac{\mu_3}{\mu_2^{3/2}} = \frac{Ex^3 - 3\mu Ex^2 + 2\mu^3}{(\sigma^2)^{3/2}}$$

$$sk = \frac{\left\{ ae^a (\alpha b)^{\frac{3}{\beta}} \tau_u \left( a, \frac{3}{\beta}, 1 \right) - 3 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\} \right\}}{\left\{ ae^a (\alpha b)^{\frac{2}{\beta}} \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} + 2 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\}^3} \quad (26)$$

$$kr = \frac{\mu_4}{\mu_2^2} - 3 = \frac{Ex^4 - 4\mu Ex^3 + 6\mu^2 Ex^2 - 3\mu^4}{(\sigma^2)^2} - 3$$

$$kr = \frac{\left\{ ae^a (\alpha b)^{\frac{4}{\beta}} \tau_u \left( a, \frac{4}{\beta}, 1 \right) - 4 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\} \right\}}{\left\{ ae^a (\alpha b)^{\frac{3}{\beta}} \tau_u \left( a, \frac{3}{\beta}, 1 \right) \right\} + 6 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\}^2} \quad (26)$$

$$kr = \frac{\left\{ ae^a (\alpha b)^{\frac{2}{\beta}} \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} - 3 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\}^4}{\left\{ ae^a (\alpha b)^{\frac{2}{\beta}} \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\}^2 - 3 \left\{ ae^a (\alpha b)^{\frac{1}{\beta}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\}^2} - 3$$

$$\begin{aligned}
& = \frac{a^4 e^{4a} (\alpha b)^{\frac{4}{\beta}} \left\{ \begin{array}{l} \frac{1}{a^3 e^{3a}} \tau_u \left( a, \frac{4}{\beta}, 1 \right) - \frac{4}{a^2 e^{2a}} \left\{ \tau_u \left( a, \frac{1}{\beta}, 1 \right) \right\} \\ \left\{ \tau_u \left( a, \frac{3}{\beta}, 1 \right) \right\} + \frac{6}{a e^a} \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} \\ - 3 \tau_u^4 \left( a, \frac{1}{\beta}, 1 \right) \end{array} \right\}}{a^4 e^{4a} (\alpha b)^{\frac{4}{\beta}} \left\{ \begin{array}{l} \frac{1}{a e^a} \tau_u \left( a, \frac{2}{\beta}, 1 \right) - \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \\ \left\{ \frac{1}{a^3 e^{3a}} \tau_u \left( a, \frac{4}{\beta}, 1 \right) - \frac{4}{a^2 e^{2a}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{3}{\beta}, 1 \right) \right\} \right\} \\ + \frac{6}{a e^a} \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} - 3 \tau_u^4 \left( a, \frac{1}{\beta}, 1 \right) \end{array} \right\}} - 3 \\
kr & = \frac{\left\{ \begin{array}{l} \frac{1}{a e^a} \tau_u \left( a, \frac{2}{\beta}, 1 \right) - \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \\ \left\{ \frac{1}{a^3 e^{3a}} \tau_u \left( a, \frac{4}{\beta}, 1 \right) - \frac{4}{a^2 e^{2a}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{3}{\beta}, 1 \right) \right\} \right\} \\ + \frac{6}{a e^a} \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} - 3 \tau_u^4 \left( a, \frac{1}{\beta}, 1 \right) \end{array} \right\}}{\left\{ \begin{array}{l} \frac{1}{a e^a} \tau_u \left( a, \frac{2}{\beta}, 1 \right) - \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \\ \left\{ \frac{1}{a^3 e^{3a}} \tau_u \left( a, \frac{4}{\beta}, 1 \right) - \frac{4}{a^2 e^{2a}} \tau_u \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{3}{\beta}, 1 \right) \right\} \right\} \\ + \frac{6}{a e^a} \tau_u^2 \left( a, \frac{1}{\beta}, 1 \right) \left\{ \tau_u \left( a, \frac{2}{\beta}, 1 \right) \right\} - 3 \tau_u^4 \left( a, \frac{1}{\beta}, 1 \right) \end{array} \right\}} - 3. \tag{27}
\end{aligned}$$

Since  $(x \leq x_q) = F_x(x_q) = \frac{e^{-ae^{\alpha bx^{-\beta}}}}{e^{-a}}$   $0 < q < 1$   $x_q > 0$ , then the quintile function  $x_q$  of [0, 1] TFF random variable can be defined as

$$x_q = F^{-1}(q) = \left[ \frac{\ln \left( 1 - \frac{\ln(q)}{a} \right)}{ab} \right]^{\frac{-1}{\beta}}$$

So, by using the inverse transform method, we can generate [0, 1] TFF random variable as follows:

$$x = \left[ \frac{\ln \left( 1 - \frac{\ln(u)}{a} \right)}{ab} \right]^{\frac{-1}{\beta}},$$

where  $u$  is a uniformly distributed random number lie in the unit interval [0,1].

### 3.1. Shannon and Relative Entropies

The Shannon entropy of [0, 1] TFF( $a, b, \alpha, \beta$ ) random variable X can be found as follows:

$$\begin{aligned}
H & = - \int_0^\infty f(x) \ln(f(x)) dx \\
& = - \int_0^\infty f(x) \ln \left( \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} \right) dx \\
& = \int_0^\infty f(x) \left[ -\ln \left( \frac{ab\alpha\beta}{e^{-a}} \right) + (\beta+1) \ln(x) - \alpha bx^{-\beta} + ae^{\alpha bx^{-\beta}} \right] dx \\
& = -\ln \left( \frac{ab\alpha\beta}{e^{-a}} \right) + (\beta+1)E(\ln(x)) - \alpha b E(x^{-\beta}) + a E(e^{\alpha bx^{-\beta}}).
\end{aligned}$$

Let  $I_1 = (\beta+1)E(\ln(x))$

$$= (\beta + 1) \int_0^\infty \ln(x) \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} \left(e^{-\alpha x^{-\beta}}\right)^{-b} e^{-a(e^{-\alpha x^{-\beta}})^{-b}} dx.$$

Since  $e^{-a(e^{-\alpha x^{-\beta}})^{-b}} = \sum_{u=0}^{\infty} \frac{(-1)^u}{u!} a^u \left(e^{-\alpha x^{-\beta}}\right)^{-bu}$ , then

$$\begin{aligned} I_1 &= (\beta + 1) \frac{ab\alpha\beta}{e^{-a}} \sum_{u=0}^{\infty} \frac{(-1)^u}{u!} a^u \int_0^\infty \ln(x) x^{-(\beta+1)} \left(e^{-\alpha x^{-\beta}}\right)^{-b} \left(e^{-\alpha x^{-\beta}}\right)^{-bu} dx \\ &= (\beta + 1) \frac{ab\alpha\beta}{e^{-a}} \sum_{u=0}^{\infty} \frac{(-1)^u}{u!} a^u \int_0^\infty \ln(x) x^{-(\beta+1)} \left(e^{-\alpha x^{-\beta}}\right)^{-b(u+1)} dx. \end{aligned}$$

Since  $\left(e^{-\alpha x^{-\beta}}\right)^{-b(u+1)} = \left(1 - (1 - e^{-\alpha x^{-\beta}})\right)^{-b(u+1)}$ , by using

$$(1-z)^{-k} = \sum_{j=0}^{\infty} \frac{\Gamma(k+j)}{j! \Gamma(k)} z^j \quad \text{and} \quad (1-z)^\delta = \sum_{s=0}^{\infty} \frac{(-1)^s}{s!} \frac{\Gamma(\delta+1)}{\Gamma(\delta-s+1)} |z| < 1, \quad k, \delta > 0$$

is non-integer real number. We get,

$$\begin{aligned} \left(1 - (1 - e^{-\alpha x^{-\beta}})\right)^{-b(u+1)} &= \sum_{j=0}^{\infty} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \sum_{s=0}^j \frac{(-1)^s}{s!} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} e^{-\alpha s x^{-\beta}}, \text{ and then} \\ I_1 &= (\beta + 1) \frac{ab\alpha\beta}{e^{-a}} \sum_{u=0}^{\infty} \frac{(-1)^{u+s}}{u! s!} a^u \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \int_0^\infty \ln(x) x^{-(\beta+1)} e^{-\alpha s x^{-\beta}} dx. \end{aligned}$$

$$\text{Let } y = x^{-\beta} \Rightarrow x = y^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} y^{\frac{-1}{\beta}-1} dy$$

$$\begin{aligned} I_1 &= (\beta + 1) \frac{ab\alpha\beta}{e^{-a}} \sum_{u=0}^{\infty} \frac{(-1)^{u+s}}{u! s!} a^u \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \int_0^\infty \ln\left(y^{\frac{-1}{\beta}}\right) y^{1+\frac{1}{\beta}} e^{-\alpha s y} \\ &\quad \frac{1}{\beta} y^{\frac{-1}{\beta}-1} dy \\ &= -(\beta + 1) \frac{bae^a}{\beta} \sum_{u=0}^{\infty} \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \int_0^\infty \ln(y) e^{-\alpha s y} dy. \end{aligned}$$

Since  $\int_0^\infty x^{s-1} e^{-mx} \ln(x) dx = m^{-s} \Gamma(s) \{\Psi(s) - \ln(m)\}$ , where  $\Psi(1) = -\gamma$  and  $s = 1, m = \alpha s$

$$\begin{aligned} I_1 &= -(\beta + 1) \frac{bae^a}{\beta} \sum_{u=0}^{\infty} \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \frac{1}{\alpha s} \{-\gamma - \ln(\alpha s)\} \\ &= (\beta + 1) \frac{be^a}{\beta s} \sum_{u=0}^{\infty} \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \{\gamma + \ln(\alpha s)\}, \end{aligned}$$

$$\text{and } I_2 = -\alpha b E(x^{-\beta}) = -\alpha b \int_0^\infty x^{-\beta} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx$$

$$= -\frac{a\beta(\alpha b)^2}{e^{-a}} \int_0^\infty x^{-2\beta-1} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx.$$

$$\text{Let } y = e^{\alpha bx^{-\beta}} \Rightarrow x = \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}-1} \frac{1}{aby} dy$$

$$\begin{aligned}
I_2 &= \frac{-a\beta(\alpha b)^2}{e^{-a}} \int_1^\infty \left( \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}} \right)^{-2\beta-1} y e^{-ay} \frac{1}{\beta} \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}-1} \frac{1}{\alpha b y} dy \\
&= -ae^a \int_1^\infty \ln(y) e^{-ay} dy \\
&= -ae^a \left[ \int_0^\infty \ln(y) e^{-ay} dy - \int_0^1 \ln(y) e^{-ay} dy \right] \\
I_{21} &= \int_0^\infty \ln(y) e^{-ay} dy = \frac{1}{a} \{-\gamma - \ln(a)\} = \frac{-1}{a} \{\gamma + \ln(a)\}.
\end{aligned}$$

Since  $\int_0^\infty x^{s-1} e^{-mx} \ln(x) dx = m^{-s} \Gamma(s) \{\Psi(s) - \ln(m)\}$ , where  $\Psi(1) = -\gamma$  and  $s = 1, m = a$

$$\begin{aligned}
I_{22} &= \int_0^1 \ln(y) e^{-ay} dy \\
&= \sum_{m=0}^{\infty} \frac{(-a)^m}{m!} \int_0^1 \ln(y) y^m dy = \sum_{m=0}^{\infty} \frac{(-1)^{m+1} a^m}{m! (m+1)^2}.
\end{aligned}$$

Since  $e^{-ay} = \sum_{m=0}^{\infty} \frac{(-ay)^m}{m!}$  and  $\int x^m \ln(x) dx = x^{m+1} \left\{ \frac{\ln(x)}{m+1} - \frac{1}{(m+1)^2} \right\}$ ,  $I_2 = e^a \left\{ \gamma + \ln(a) + \sum_{m=0}^{\infty} \frac{(-1)^{m+1} a^{m+1}}{m! (m+1)^2} \right\}$ , and  $I_3 = aE(e^{\alpha bx - \beta})$

$$= a \int_0^\infty e^{\alpha bx - \beta} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx - \beta} e^{-ae^{\alpha bx - \beta}} dx.$$

Let  $y = e^{\alpha bx - \beta} \Rightarrow x = \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}-1} \frac{1}{\alpha b} \frac{1}{y} dy$ , then

$$\begin{aligned}
I_3 &= \frac{a^2 b \alpha \beta}{e^{-a}} \int_1^\infty \left( \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}} \right)^{-(\beta+1)} y^2 e^{-ay} \frac{1}{\beta} \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}-1} \frac{1}{\alpha b} \frac{1}{y} dy \\
&= a^2 e^a \int_1^\infty \left( \frac{\ln y}{\alpha b} \right)^{1+\frac{1}{\beta}} y e^{-ay} \left( \frac{\ln y}{\alpha b} \right)^{\frac{-1}{\beta}-1} dy \\
I_3 &= a^2 e^a \int_1^\infty y e^{-ay} dy = a + 1
\end{aligned}$$

$$\begin{aligned}
H &= \ln \left( \frac{e^{-a}}{ab\alpha\beta} \right) + \frac{(\beta+1)be^a}{\beta s} \sum_{u=0}^{\infty} \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \{\gamma \\
&\quad + \ln(as)\} + e^a \left\{ \gamma + \ln(a) + \sum_{m=0}^{\infty} \frac{(-1)^{m+1} a^{m+1}}{m! (m+1)^2} \right\} + a + 1. \tag{28}
\end{aligned}$$

The relative entropy  $Dkl(F||F^*)$  for a random variable  $[0, 1]$  TFF( $a, b, \alpha, \beta$ ) can be found as follows:

$$\begin{aligned}
\frac{f(x)}{f^*(x)} &= \frac{\frac{ab\alpha\beta}{e^{-a}}x^{-(\beta+1)}e^{\alpha bx^{-\beta}}e^{-ae^{\alpha bx^{-\beta}}}}{\frac{a_1b_1\alpha_1\beta_1}{e^{-a_1}}x^{-(\beta_1+1)}e^{\alpha_1 b_1 x^{-\beta_1}}e^{-a_1 e^{\alpha_1 b_1 x^{-\beta_1}}}} \\
Dkl(F_1 \| F_1^*) &= \int_0^\infty f(x) \ln \left( \frac{ab\alpha\beta e^{-a_1}x^{-(\beta+1)}e^{\alpha bx^{-\beta}}e^{-ae^{\alpha bx^{-\beta}}}}{a_1b_1\alpha_1\beta_1 e^{-a}x^{-(\beta_1+1)}e^{\alpha_1 b_1 x^{-\beta_1}}e^{-a_1 e^{\alpha_1 b_1 x^{-\beta_1}}}} \right) dx \\
&= \int_0^\infty f_1(x) \left[ \ln \left( \frac{ab\alpha\beta e^{-a_1}}{a_1b_1\alpha_1\beta_1 e^{-a}} \right) + (\beta_1 - \beta) \ln(x) + \alpha bx^{-\beta} - ae^{\alpha bx^{-\beta}} - \frac{a_1b_1 x^{-\beta_1} + a_1 e^{\alpha_1 b_1 x^{-\beta_1}}}{a_1b_1\alpha_1\beta_1 e^{-a}} \right] dx \\
&= \ln \left( \frac{ab\alpha\beta e^{-a_1}}{a_1b_1\alpha_1\beta_1 e^{-a}} \right) + (\beta_1 - \beta)E(\ln(x)) + \alpha b E(x^{-\beta}) - a E(e^{\alpha bx^{-\beta}}) - a_1 b_1 E(x^{-\beta_1}) \\
&\quad + a_1 E(e^{\alpha_1 b_1 x^{-\beta_1}}).
\end{aligned}$$

Let,  $I_1 = (\beta_1 - \beta)E(\ln(x))$

$$\begin{aligned}
&= (\beta_1 - \beta) \frac{ab\alpha\beta}{e^{-a}} \int_0^\infty \ln(x) x^{-(\beta+1)} (e^{-\alpha x^{-\beta}})^{-b} e^{-a(e^{-\alpha x^{-\beta}})^{-b}} dx \\
I_1 &= \frac{(\beta_1 - \beta)be^a}{\beta s} \sum_{u=0}^\infty \sum_{j=0}^\infty \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \{Y + \ln(as)\},
\end{aligned}$$

and  $I_2 = \alpha b E(x^{-\beta})$

$$\begin{aligned}
&= \alpha b \int_0^\infty x^{-\beta} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx \\
I_2 &= -e^a \left\{ Y + \ln(a) + \sum_{m=0}^\infty \frac{(-1)^{m+1} a^{m+1}}{m! (m+1)^2} \right\},
\end{aligned}$$

and  $I_3 = -a E(e^{\alpha bx^{-\beta}})$

$$\begin{aligned}
&= -a \int_0^\infty e^{\alpha bx^{-\beta}} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx \\
I_3 &= -(a+1),
\end{aligned}$$

and  $I_4 = -\alpha_1 b_1 E(x^{-\beta_1})$

$$= -\alpha_1 b_1 \int_0^\infty x^{-\beta_1} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx^{-\beta}} e^{-ae^{\alpha bx^{-\beta}}} dx.$$

Let  $y = e^{\alpha bx^{-\beta}} \Rightarrow x = \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}-1} \frac{1}{\alpha b} \frac{1}{y} dy$

$$\begin{aligned}
I_4 &= \frac{-\alpha_1 b_1 ab\alpha\beta}{e^{-a}} \int_1^\infty \left( \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}} \right)^{-(\beta_1+\beta+1)} y e^{-ay} \frac{1}{\beta} \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}-1} \frac{1}{\alpha b} \frac{1}{y} dy \\
&= \frac{-\alpha_1 b_1 a}{e^{-a}} \int_1^\infty \left(\frac{\ln y}{\alpha b}\right)^{\frac{\beta_1+1+\frac{1}{\beta}}{\beta}} e^{-ay} \left(\frac{\ln y}{\alpha b}\right)^{\frac{-1}{\beta}-1} dy
\end{aligned}$$

$$= -\alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \int_1^\infty (\ln y)^{\frac{\beta_1}{\beta}} e^{-ay} dy.$$

Since  $e^{-ay} = \sum_{m=0}^\infty \frac{(-a)^m}{m!} y^m$ , then

$$I_4 = -\alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-a)^m}{m!} \int_1^\infty (\ln y)^{\frac{\beta_1}{\beta}} y^m dy.$$

Let  $y = e^{-t^{-1}} \Rightarrow t = (-\ln y)^{-1} \Rightarrow dy = t^{-2} e^{-t^{-1}} dt$ , then

$$\begin{aligned} I_4 &= \alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-a)^m}{m!} \int_0^\infty (\ln e^{-t^{-1}})^{\frac{\beta_1}{\beta}} (e^{-t^{-1}})^m t^{-2} e^{-t^{-1}} dt \\ &= \alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-a)^m}{m!} (-1)^{\frac{\beta_1}{\beta}} \int_0^\infty t^{\frac{-\beta_1}{\beta}} t^{-2} e^{-(m+1)t^{-1}} dt. \end{aligned}$$

Let  $w = t^{-1} \Rightarrow t = w^{-1} \Rightarrow dt = -w^{-2} dw$

$$\begin{aligned} I_4 &= \alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-a)^m}{m!} (-1)^{\frac{\beta_1}{\beta}} \int_0^\infty w^{\frac{\beta_1}{\beta}} w^2 e^{-(m+1)w} w^{-2} dw \\ &= \alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-a)^m}{m!} (-1)^{\frac{\beta_1}{\beta}} \int_0^\infty w^{\frac{\beta_1}{\beta}} e^{-(m+1)w} dw \\ &= \alpha_1 b_1 a e^a (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^\infty \frac{(-1)^{m+\frac{\beta_1}{\beta}}}{m! (m+1)^{\frac{\beta_1}{\beta}}} a^{m+1} \Gamma\left(\frac{\beta_1}{\beta} + 1\right), \end{aligned}$$

and  $I_5 = a_1 E\left(e^{\alpha_1 b_1 x^{-\beta_1}}\right)$

$$= \frac{a_1 ab \alpha \beta}{e^{-a}} \int_0^\infty e^{\alpha_1 b_1 x^{-\beta_1}} x^{-(\beta+1)} e^{\alpha b x^{-\beta}} e^{-a e^{\alpha b x^{-\beta}}} dx.$$

Since  $e^{\alpha_1 b_1 x^{-\beta_1}} = \sum_{s=0}^\infty \frac{(\alpha_1 b_1)^s}{s!} x^{-\beta_1 s}$ , then

$$I_5 = \frac{a_1 ab \alpha \beta}{e^{-a}} \sum_{s=0}^\infty \frac{(\alpha_1 b_1)^s}{s!} \int_0^\infty x^{-\beta_1 s - \beta - 1} e^{\alpha b x^{-\beta}} e^{-a e^{\alpha b x^{-\beta}}} dx.$$

Let  $y = e^{\alpha b x^{-\beta}} \Rightarrow x = \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} \frac{1}{aby} dy$

$$\begin{aligned} I_5 &= \frac{a_1 ab \alpha \beta}{e^{-a}} \sum_{s=0}^\infty \frac{(\alpha_1 b_1)^s}{s!} \int_1^\infty \left(\frac{\ln y}{ab}\right)^{-\beta_1 s - \beta - 1} y e^{-ay} \frac{1}{\beta} \left(\frac{\ln y}{ab}\right)^{\frac{-1}{\beta}-1} \frac{1}{aby} dy \\ &= a_1 a e^a \sum_{s=0}^\infty \frac{(\alpha_1 b_1)^s}{s!} \int_1^\infty \left(\frac{\ln y}{ab}\right)^{\frac{\beta_1 s}{\beta}} e^{-ay} dy. \end{aligned}$$

So, by using,  $e^{-ay} = \sum_{m=0}^\infty \frac{(-a)^m}{m!} y^m$ , we get

$$I_5 = a_1 a e^a \sum_{s=0}^\infty \frac{(\alpha_1 b_1)^s}{s!} \sum_{m=0}^\infty \frac{(-a)^m}{m!} \int_1^\infty \left(\frac{\ln y}{ab}\right)^{\frac{\beta_1 s}{\beta}} y^m dy.$$

Let  $y = e^{-t^{-1}} \Rightarrow t = (-\ln y)^{-1} \Rightarrow dy = t^{-2}e^{-t^{-1}}dt$

$$\begin{aligned} I_5 &= -a_1ae^a \sum_{s=0}^{\infty} \frac{(\alpha_1 b_1)^s}{s!} \sum_{m=0}^{\infty} \frac{(-a)^m}{m!} \int_0^{\infty} \left( \frac{\ln e^{-t^{-1}}}{ab} \right)^{\frac{\beta_1 s}{\beta}} (e^{-t^{-1}})^m t^{-2} e^{-t^{-1}} dt \\ &= -a_1e^a \sum_{s=0}^{\infty} \frac{(\alpha_1 b_1)^s}{s!} \sum_{m=0}^{\infty} \frac{(-1)^m}{m!} a^{m+1} (-1)^{\frac{\beta_1 s}{\beta}} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \int_0^{\infty} t^{\frac{-\beta_1 s}{\beta}} e^{-(m+1)t^{-1}} t^{-2} dt. \end{aligned}$$

Let  $w = t^{-1} \Rightarrow t = w^{-1} \Rightarrow dt = -w^{-2}dw$

$$\begin{aligned} I_5 &= -a_1e^a \sum_{s,m=0}^{\infty} \frac{(\alpha_1 b_1)^s}{s!} \frac{(-1)^{m+\frac{\beta_1 s}{\beta}}}{m!} a^{m+1} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \int_0^{\infty} w^{\frac{\beta_1 s}{\beta}} e^{-(m+1)w} dw \\ &= -a_1e^a \sum_{s,m=0}^{\infty} \frac{(\alpha_1 b_1)^s}{s!} \frac{(-1)^{m+\frac{\beta_1 s}{\beta}}}{m!} a^{m+1} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \frac{\Gamma\left(\frac{\beta_1 s}{\beta} + 1\right)}{(m+1)^{\frac{\beta_1 s}{\beta} + 1}} \end{aligned}$$

$$\begin{aligned} Dkl(F_1 \| F_1^*) &= \ln \left( \frac{ab\alpha\beta e^{-a_1}}{a_1 b_1 \alpha_1 \beta_1 e^{-a}} \right) + (\beta_1 - \beta) \frac{be^a}{\beta s} \sum_{u=0}^{\infty} \sum_{j=0}^{\infty} \sum_{s=0}^j \frac{(-1)^{u+s}}{u! s!} a^{u+1} \frac{\Gamma(b(u+1)+j)}{j! \Gamma(b(u+1))} \\ &\quad \frac{\Gamma(j+1)}{\Gamma(j-s+1)} \{Y + \ln(as)\} - e^a \left\{ Y + \ln(a) + \sum_{m=0}^{\infty} \frac{(-1)^{m+1} a^{m+1}}{m! (m+1)^2} \right\} - (a+1) + \alpha_1 b_1 e^a \\ &\quad (\alpha b)^{\frac{-\beta_1}{\beta}} \sum_{m=0}^{\infty} \frac{(-1)^{m+\frac{\beta_1}{\beta}}}{m!} a^{m+1} \frac{\Gamma\left(\frac{\beta_1}{\beta} + 1\right)}{(m+1)^{\frac{\beta_1}{\beta} + 1}} \\ &\quad - a_1e^a \sum_{s,m=0}^{\infty} \frac{(\alpha_1 b_1)^s}{s!} \frac{(-1)^{m+\frac{\beta_1 s}{\beta}}}{m!} a^{m+1} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \frac{\Gamma\left(\frac{\beta_1 s}{\beta} + 1\right)}{(m+1)^{\frac{\beta_1 s}{\beta} + 1}}. \end{aligned} \tag{29}$$

### 3.2. Stress-Strength Reliability

Let  $y$  and  $x$  be the stress and the strength random variables, independent of each other, follow  $[0, 1]$  TFF( $a, b, \alpha, \beta$ ) and  $[0, 1]$  TFF( $\alpha_1, b_1, \alpha_1, \beta_1$ ), respectively. Then

$$\begin{aligned} R &= P(Y < X) = \int_0^{\infty} f_x(x) F_y(x) dx \\ &= \int_0^{\infty} \frac{ab\alpha\beta}{e^{-a}} x^{-(\beta+1)} e^{\alpha bx - \beta} e^{-ae^{\alpha bx} - \beta} \frac{e^{-a_1 e^{\alpha_1 b_1 x} - \beta_1}}{e^{-a_1}} dx. \end{aligned}$$

Since  $e^{-a_1 e^{\alpha_1 b_1 x} - \beta_1} = \sum_{i=0}^{\infty} \sum_{s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} x^{-\beta_1 s}$ , then

$$R = \frac{ab\alpha\beta}{e^{-(a+a_1)}} \sum_{i,s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} \int_0^{\infty} x^{-\beta_1 s - \beta - 1} e^{\alpha bx - \beta} e^{-ae^{\alpha bx} - \beta} dx.$$

$$\text{Let } y = e^{\alpha bx^{-\beta}} \Rightarrow x = \left(\frac{\ln y}{ab}\right)^{-\frac{1}{\beta}} \Rightarrow dx = \frac{-1}{\beta} \left(\frac{\ln y}{ab}\right)^{-\frac{1}{\beta}-1} \frac{1}{ab} \frac{1}{y} dy$$

$$R = \frac{ab\alpha\beta}{e^{-(a+a_1)}} \sum_{i,s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} \int_1^{\infty} \left(\frac{\ln y}{ab}\right)^{\frac{\beta_1 s+1+\beta}{\beta}} y e^{-ay} \frac{1}{\beta} \left(\frac{\ln y}{ab}\right)^{-\frac{1}{\beta}-1} \frac{1}{ab} \frac{1}{y} dy$$

$$= ae^{a+a_1} \sum_{i,s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} \int_1^{\infty} \left(\frac{\ln y}{ab}\right)^{\frac{\beta_1 s}{\beta}} e^{-ay} dy.$$

So, by using  $e^{-ay} = \sum_{m=0}^{\infty} \frac{(-a)^m}{m!} y^m$ , we get

$$R = ae^{a+a_1} \sum_{i,s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} \sum_{m=0}^{\infty} \frac{(-a)^m}{m!} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \int_1^{\infty} (\ln y)^{\frac{\beta_1 s}{\beta}} y^m dy.$$

Let  $y = e^{-t^{-1}} \Rightarrow t = (-\ln y)^{-1} \Rightarrow dy = t^{-2} e^{-t^{-1}} dt$

$$R = -ae^{a+a_1} \sum_{i,s=0}^{\infty} \frac{(-a_1)^i}{i!} \frac{(i\alpha_1 b_1)^s}{s!} \sum_{m=0}^{\infty} \frac{(-a)^m}{m!} (\alpha b)^{\frac{-\beta_1 s}{\beta}} \int_0^{\infty} (\ln e^{-t^{-1}})^{\frac{\beta_1 s}{\beta}} (e^{-t^{-1}})^m t^{-2} e^{-t^{-1}} dt$$

$$= -e^{a+a_1} \sum_{i,s,m=0}^{\infty} \frac{(-1)^{i+m+\frac{\beta_1 s}{\beta}}}{i! m!} a_1^i \frac{(i\alpha_1 b_1 (\alpha b)^{-\beta_1/\beta})^s}{s!} a^{m+1} \int_0^{\infty} t^{\frac{-\beta_1 s}{\beta}} e^{-(m+1)t^{-1}} t^{-2} dt.$$

Let  $w = t^{-1} \Rightarrow t = w^{-1} \Rightarrow dt = -w^{-2} dw$

$$R = -e^{a+a_1} \sum_{i,s,m=0}^{\infty} \frac{(-1)^{i+m+\frac{\beta_1 s}{\beta}}}{i! m!} a_1^i \frac{(i\alpha_1 b_1 (\alpha b)^{-\beta_1/\beta})^s}{s!} a^{m+1} \int_0^{\infty} w^{\frac{\beta_1 s}{\beta}} e^{-(m+1)w} w^2 w^{-2} dw$$

$$= -e^{a+a_1} \sum_{i,s,m=0}^{\infty} \frac{(-1)^{i+m+\frac{\beta_1 s}{\beta}}}{i! m!} a_1^i \frac{(i\alpha_1 b_1 (\alpha b)^{-\beta_1/\beta})^s}{s!} a^{m+1} \int_0^{\infty} w^{\frac{\beta_1 s}{\beta}} e^{-(m+1)w} dw$$

$$R = -e^{a+a_1} \sum_{i,s,m=0}^{\infty} \frac{(-1)^{i+m+\frac{\beta_1 s}{\beta}}}{i! m!} a_1^i \frac{(i\alpha_1 b_1 (\alpha b)^{-\beta_1/\beta})^s}{s!} a^{m+1} \frac{\Gamma\left(\frac{\beta_1 s}{\beta} + 1\right)}{(m+1)^{\frac{\beta_1 s}{\beta} + 1}}.$$

#### 4. Summary and Conclusions

In statistical analysis, a lot of distributions are used to represent set(s) data. Recently, new distributions are derived to extend some of well-known families of distributions; such the new distributions are more flexible than the others to model real data are. The composing of some distributions with each other's some way has been in the foreword of data modeling. In this paper, we presented a new family of continuous distributions based on [0, 1] truncated Fréchet distribution. [0, 1] Truncated Fréchet-Weibull ([0, 1] TFU ) and [0, 1] Truncated Fréchet- Fréchet ([0,1 ] TFF ) distributions that were discussed as special cases. Properties of [0, 1] TFW and [0, 1] TFF was derived. We provided forms for characteristic function, rth raw moment, mean, variance, skewness, kurtosis, median, reliability function,

hazard rate function, Shannon entropy function, and Relative entropy function. This paper also dealt with the determination of stress strength reliability  $R = P[Y < X]$  when X (strength) and Y (stress) were two independent [0, 1] TFW ([0, 1] TFF) distributions with different parameters.

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